

## CHAPTER 3

### Main Results

#### 3.1 Measure of Complete Dependence Based on Conditional Distribution Functions

In this chapter, we define the measure of dependence for random vectors using the conditional distribution functions.

**Definition 3.1.1.** Let  $X$  and  $Y$  be random vectors.

The *measure of dependence*  $\varphi$  of  $Y$  given  $X$  is defined by

$$\varphi(Y|X) = \int \int |F_{Y|X}(v|u) - F_Y(v)| dF_X(u) dF_Y(v)$$

where  $F_{Y|X}$  is the conditional distribution function of  $Y$  given  $X$ .

We will figure out the maximum value of  $\varphi$  using Lemma 3.1.3. The prove of Lemma 3.1.3 is very complicated. Therefore, we separated parts of its proof into the following lemmas.

**Lemma 3.1.1.** Let  $A$  be a metric space and  $\mu$  be a Borel probability measure on  $A$  such that  $\mu(\overline{\mathcal{B}(x, \epsilon)} \setminus \mathcal{B}(x, \epsilon)) = 0$  for all ball  $\mathcal{B}(x, \epsilon)$  centered in  $x \in A$  and of radius  $\epsilon > 0$ . Let  $y \in (0, 1)$  and  $\mathfrak{D}_y = \{f : A \rightarrow [0, 1] \mid f \text{ is measurable and } \int_A f(x) d\mu = y\}$ . Then a function  $H_1$  defined by  $H_1(\epsilon) = \int_{\mathcal{B}(x, \epsilon) \cap E_f^c} f(z) d\mu - \int_{\mathcal{B}^c(x, \epsilon) \cap E_f^c} (1 - f(z)) d\mu$  is a continuous function where  $E_f = \{x \in A \mid f(x) > y\}$  and  $f \in \mathfrak{D}_y$ .

*Proof.* Let  $y \in \mathfrak{D}_y$ . Consider

$$\begin{aligned}
H_1(a) &= \int_{\mathcal{B}(x,a) \cap E_f^c} f(z) d\mu - \int_{\mathcal{B}^c(x,a) \cap E_f^c} (1 - f(z)) d\mu \\
&= \int_{\mathcal{B}(x,a) \cap E_f^c} f(z) d\mu - \int_{\mathcal{B}^c(x,a) \cap E_f^c} 1 d\mu + \int_{\mathcal{B}^c(x,a) \cap E_f^c} f(z) d\mu \\
&= \int_{\mathcal{B}(x,a) \cap E_f^c} f(z) d\mu + \int_{\mathcal{B}^c(x,a) \cap E_f^c} f(z) d\mu - \int_{\mathcal{B}^c(x,a) \cap E_f^c} 1 d\mu \\
&= \int_{E_f^c} f(z) d\mu - \mu(\mathcal{B}^c(x,a) \cap E_f^c) \\
&= \int_{E_f^c} f(z) d\mu - (1 - \mu(\mathcal{B}(x,a)) + \mu(E_f^c) - \mu(\mathcal{B}^c(x,a) \cup E_f^c)) \\
&= \int_{E_f^c} f(z) d\mu - 1 + \mu(\mathcal{B}(x,a)) - \mu(E_f^c) + \mu(\mathcal{B}^c(x,a) \cup E_f^c) \\
&= \int_{E_f^c} f(z) d\mu + \mu(\mathcal{B}(x,a)) - \mu(E_f^c) + (-1 + \mu(\mathcal{B}^c(x,a) \cup E_f^c)) \\
&= \int_{E_f^c} f(z) d\mu + \mu(\mathcal{B}(x,a)) - \mu(E_f^c) - (1 - \mu(\mathcal{B}^c(x,a) \cup E_f^c)) \\
&= \int_{E_f^c} f(z) d\mu + \mu(\mathcal{B}(x,a)) - \mu(E_f^c) - \mu(\mathcal{B}(x,a) \cap E_f).
\end{aligned}$$

Let  $b_n \nearrow a$ .

Then  $\mathcal{B}(x, b_n) \subseteq \mathcal{B}(x, a)$  and  $\bigcup_n \mathcal{B}(x, b_n) = \mathcal{B}(x, a)$ .

So,  $\mu(\mathcal{B}(x, b_n)) \rightarrow \mu(\mathcal{B}(x, a))$  and  $\mu(\mathcal{B}(x, b_n) \cap E_f) \rightarrow \mu(\mathcal{B}(x, a) \cap E_f)$ .

Thus,  $\lim_{b_n \rightarrow a^-} H_1(b_n) = H_1(a)$ .

If  $b_n \searrow a$ , then  $\bigcap_n \mathcal{B}(x, b_n) = \overline{\mathcal{B}(x, a)}$  and  $\mu(\overline{\mathcal{B}(x, a)} \setminus \mathcal{B}(x, a)) = 0$ , we can conclude that  $\mu(\mathcal{B}(x, b_n)) \searrow \mu(\mathcal{B}(x, a))$  and  $\mu(\mathcal{B}(x, b_n) \cap E_f) \searrow \mu(\mathcal{B}(x, a) \cap E_f)$ .

Thus,  $\lim_{b_n \rightarrow a^+} H_1(b_n) = H_1(a)$ .

Hence,  $H_1$  is a continuous function. □

**Lemma 3.1.2.** Let  $A$  be a metric space and  $\mu$  be a Borel probability measure on  $A$  such that  $\mu(\overline{\mathcal{B}(x, \epsilon)} \setminus \mathcal{B}(x, \epsilon)) = 0$  for all ball  $\mathcal{B}(x, \epsilon)$  centered in  $x \in A$  and of radius  $\epsilon > 0$ . Let  $y \in (0, 1)$  and  $\mathfrak{D}_y = \{f : A \rightarrow [0, 1] \mid f \text{ is measurable and } \int_A f(x) d\mu = y\}$ . Then a function  $H_2$  defined by  $H_2(\epsilon) = \int_{\mathcal{B}(x, \epsilon) \cap E_f} f(z) d\mu - \int_{\mathcal{B}^c(x, \epsilon) \cap E_f} (1 - f(z)) d\mu$  is a continuous function where  $E_f = \{x \in A \mid f(x) > y\}$  and  $f \in \mathfrak{D}_y$ .

*Proof.* Let  $y \in \mathfrak{D}_y$ . Consider

$$\begin{aligned}
H_2(a) &= \int_{\mathcal{B}(x, a) \cap E_f} f(z) d\mu - \int_{\mathcal{B}^c(x, a) \cap E_f} (1 - f(z)) d\mu \\
&= \int_{\mathcal{B}(x, a) \cap E_f} f(z) d\mu - \int_{\mathcal{B}^c(x, a) \cap E_f} 1 d\mu + \int_{\mathcal{B}^c(x, a) \cap E_f} f(z) d\mu \\
&= \int_{\mathcal{B}(x, a) \cap E_f} f(z) d\mu + \int_{\mathcal{B}^c(x, a) \cap E_f} f(z) d\mu - \int_{\mathcal{B}^c(x, a) \cap E_f} 1 d\mu \\
&= \int_{E_f} f(z) d\mu - \mu(\mathcal{B}^c(x, a) \cap E_f) \\
&= \int_{E_f} f(z) d\mu - (1 - \mu(\mathcal{B}(x, a))) + \mu(E_f) - \mu(\mathcal{B}^c(x, a) \cup E_f) \\
&= \int_{E_f} f(z) d\mu - 1 + \mu(\mathcal{B}(x, a)) - \mu(E_f) + \mu(\mathcal{B}^c(x, a) \cup E_f) \\
&= \int_{E_f} f(z) d\mu + \mu(\mathcal{B}(x, a)) - \mu(E_f) + (-1 + \mu(\mathcal{B}^c(x, a) \cup E_f)) \\
&= \int_{E_f} f(z) d\mu + \mu(\mathcal{B}(x, a)) - \mu(E_f) - (1 - \mu(\mathcal{B}^c(x, a) \cup E_f)) \\
&= \int_{E_f} f(z) d\mu + \mu(\mathcal{B}(x, a)) - \mu(E_f) - \mu(\mathcal{B}(x, a) \cap E_f^c).
\end{aligned}$$

Let  $b_n \nearrow a$ .

Then  $\mathcal{B}(x, b_n) \subseteq \mathcal{B}(x, a)$  and  $\bigcup_n \mathcal{B}(x, b_n) = \mathcal{B}(x, a)$ .

So,  $\mu(\mathcal{B}(x, b_n)) \xrightarrow{n} \mu(\mathcal{B}(x, a))$  and  $\mu(\mathcal{B}(x, b_n) \cap E_f^c) \rightarrow \mu(\mathcal{B}(x, a) \cap E_f^c)$ .

Thus,  $\lim_{b_n \rightarrow a^-} H_2(b_n) = H_2(a)$ .

If  $b_n \searrow a$ , then  $\bigcap_n \mathcal{B}(x, b_n) = \overline{\mathcal{B}(x, a)}$  and  $\mu(\overline{\mathcal{B}(x, a)} \setminus \mathcal{B}(x, a)) = 0$ , we can conclude that  $\mu(\mathcal{B}(x, b_n)) \searrow \mu(\mathcal{B}(x, a))$  and  $\mu(\mathcal{B}(x, b_n) \cap E_f^c) \searrow \mu(\mathcal{B}(x, a) \cap E_f^c)$ .

Thus,  $\lim_{b_n \rightarrow a^+} H_2(b_n) = H_2(a)$ .

Hence,  $H_2$  is a continuous function. □

**Lemma 3.1.3.** *Let  $A$  be a metric space and  $\mu$  be a Borel probability measure on  $A$  such that  $\mu(\overline{\mathcal{B}(x, \epsilon)} \setminus \mathcal{B}(x, \epsilon)) = 0$  for all ball  $\mathcal{B}(x, \epsilon)$  centered in  $x \in A$  and of radius  $\epsilon > 0$ .*

*Let  $y \in (0, 1)$  and  $\mathfrak{D}_y = \{f : A \rightarrow [0, 1] \mid f \text{ is measurable and } \int_A f(x) d\mu = y\}$ .*

*The supremum of  $\int_A |f(x) - y| d\mu$  over  $f \in \mathfrak{D}_y$  happens when  $f \in \{0, 1\}$  a.e.*

*Moreover,  $\max_{\{f \in \mathfrak{D}_y\}} \int_A |f(x) - y| d\mu = 2y(1 - y)$ .*

*Proof.* For each  $y \in (0, 1)$ , we can find ball  $\mathcal{B}(x_0, \epsilon_0) \subseteq A$  such that  $\int_{\mathcal{B}(x_0, \epsilon_0)} d\mu = y$  by continuity of the function  $\epsilon \mapsto \mu(\mathcal{B}(x, \epsilon))$  with infimum zero and supremum one. We define a function  $f : A \rightarrow \{0, 1\}$  via

$$f(x) = \begin{cases} 1 & \text{if } x \in \mathcal{B}(x_0, \epsilon_0), \\ 0 & \text{if } x \notin \mathcal{B}(x_0, \epsilon_0). \end{cases}$$

Since  $\mu(\overline{\mathcal{B}(x, \epsilon)} \setminus \mathcal{B}(x, \epsilon)) = 0$ , we can conclude that  $f$  is an indicator function in  $\mathfrak{D}_y$ .

Let  $\mathbf{1}_B \in \mathfrak{D}_y$  be an indicator function of  $B \subseteq A$ . We shall show that

$$\int_A |\mathbf{1}_B(x) - y| d\mu = 2y(1 - y).$$

Since  $\mathbf{1}_B \in \mathfrak{D}_y$ , we can conclude that  $\mu(B) = \int \mathbf{1}_B d\mu = y$ . Consider,

$$\begin{aligned} \int_A |\mathbf{1}_B(x) - y| d\mu &= \int_B (1 - y) d\mu + \int_{B^c} y d\mu \\ &= (1 - y)\mu(B) + y\mu(B^c) \\ &= (1 - y)y + y(1 - y) \\ &= 2y(1 - y). \end{aligned} \tag{3.1}$$

Therefore,  $\int_A |\mathbf{1}_B(x) - y| d\mu = 2y(1 - y)$ .

Let  $f \in \mathfrak{D}_y$  be not an indicator function and  $E_f = \{x \in A \mid f(x) > y\}$ .

Since

$$\begin{aligned} \int_{E_f} y d\mu + \int_{E_f^c} y d\mu &= y \\ &= \int_A f(x) d\mu \\ &= \int_{E_f} f(x) d\mu + \int_{E_f^c} f(x) d\mu, \end{aligned}$$

we get

$$\int_{E_f^c} (y - f(x)) d\mu = \int_{E_f} (f(x) - y) d\mu. \tag{3.2}$$

Consider

$$\begin{aligned}
\int_A |f(x) - y| d\mu &= \int_{E_f} |f(x) - y| d\mu + \int_{E_f^c} |f(x) - y| d\mu \\
&= \int_{E_f} (f(x) - y) d\mu + \int_{E_f^c} (y - f(x)) d\mu \\
&= 2 \int_{E_f} (f(x) - y) d\mu.
\end{aligned}$$

Since  $\int_{E_f} (f(x) - y) d\mu = \int_{E_f^c} (y - f(x)) d\mu$  and  $\int_A |f(x) - y| d\mu = 2 \int_{E_f} (f(x) - y) d\mu$ , we can conclude that  $\int_A |f(x) - y| d\mu = 2 \int_{E_f^c} (y - f(x)) d\mu$ .

Next, we show that  $\int_A |f(x) - y| d\mu \leq \int_A |f^*(x) - y| d\mu$  where  $F^* \in \mathfrak{D}_y$  is an indicator function.

Case i)  $\int_{E_f^c} f(x) d\mu > 0$ .

Let  $x \in E_f$  be fixed.

Since the function  $H_1$  in Lemma 3.1.1 is continuous and  $y \geq f(x)$  for all  $x \in E_f^c$ , we get

$$\begin{aligned}
0 &> y - 1 \\
&\geq f(x) - 1 \\
&= -(1 - f(x)).
\end{aligned}$$

Then  $H_1(0) = -\int_{E_f^c} (1 - f(z)) d\mu < 0$  and  $H_1(\infty) = \int_{E_f^c} f(z) d\mu > 0$ .

Thus, there exists  $\epsilon_0 \in (0, \infty)$  such that  $H_1(\epsilon_0) = 0$ .

Define a new function  $f^*$  by  $f^* = f \mathbf{1}_{E_f} + \mathbf{1}_{E_f^c \cap \mathcal{B}^c(x, \epsilon_0)}$ .

We show that  $f^* \in \mathfrak{D}_y$ .

Since  $H_1(\epsilon_0) = 0$ , we have

$$\begin{aligned}
0 &= \int_{\mathcal{B}(x, \epsilon_0) \cap E_f^c} f(z) d\mu - \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c} (1 - f(z)) d\mu \\
&= \int_{\mathcal{B}(x, \epsilon_0) \cap E_f^c} f(z) d\mu - \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c} 1 d\mu + \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c} f(z) d\mu.
\end{aligned}$$

Therefore,

$$\int_{E_f^c} f(z) d\mu = \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c} 1 d\mu. \tag{3.3}$$

Then

$$\begin{aligned}
\int_A f^*(x) d\mu &= \int_A f \mathbf{1}_{E_f}(x) d\mu + \int_A \mathbf{1}_{B^c(x, \epsilon_0) \cap E_f^c}(x) d\mu \\
&= \int_{E_f} f(x) d\mu + \int_{B^c(x, \epsilon_0) \cap E_f^c} d\mu \\
&= \int_{E_f} f(x) d\mu + \int_{E_f^c} f(z) d\mu \\
&= \int_A f(x) d\mu \\
&= y.
\end{aligned}$$

Hence,  $f^* \in \mathfrak{D}_y$ .

Next, we show that  $\int_A |f(x) - y| d\mu < \int_A |f^*(x) - y| d\mu$ .

Since  $y \in (0, 1)$  and equation(3.3), we obtain that

$$\begin{aligned}
\int_{B(x, \epsilon_0) \cap E_f^c} y d\mu + \int_{E_f^c} f(x) d\mu &= \int_{B(x, \epsilon_0) \cap E_f^c} y d\mu + \int_{B^c(x, \epsilon_0) \cap E_f^c} d\mu \\
&> \int_{B(x, \epsilon_0) \cap E_f^c} y d\mu + \int_{B^c(x, \epsilon_0) \cap E_f^c} y d\mu \\
&= \int_{E_f^c} y d\mu.
\end{aligned}$$

Therefore,  $\int_{B(x, \epsilon_0) \cap E_f^c} y d\mu + \int_{E_f^c} f(x) d\mu > \int_{E_f^c} y d\mu$ .

Thus,

$$\int_{B(x, \epsilon_0) \cap E_f^c} y d\mu > \int_{E_f^c} (y - f(x)) d\mu. \quad (3.4)$$

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Consider

$$\begin{aligned}
\int_A |f^*(x) - y| d\mu &= \int_{E_f} |f^*(x) - y| d\mu + \int_{E_f^c} |f^*(x) - y| d\mu \\
&= \int_{E_f} |f \mathbf{1}_{E_f}(x) + \mathbf{1}_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c}(x) - y| d\mu \\
&\quad + \int_{E_f^c} |f \mathbf{1}_{E_f}(x) + \mathbf{1}_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c}(x) - y| d\mu \\
&= \int_{E_f} |f(x) - y| d\mu + \int_{E_f^c} |\mathbf{1}_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c}(x) - y| d\mu \\
&= \int_{E_f} |f(x) - y| d\mu + \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c} (1 - y) d\mu + \int_{\mathcal{B}(x, \epsilon_0) \cap E_f^c} y d\mu \\
&= \int_{E_f} f(x) d\mu - \int_{E_f} y d\mu + \int_{E_f^c} f(x) d\mu \\
&\quad - \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c} y d\mu + \int_{\mathcal{B}(x, \epsilon_0) \cap E_f^c} y d\mu \\
&= \int_A f(x) d\mu - \int_{E_f} y d\mu - \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c} y d\mu + \int_{\mathcal{B}(x, \epsilon_0) \cap E_f^c} y d\mu \\
&= y - \int_{E_f} y d\mu - \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c} y d\mu + \int_{\mathcal{B}(x, \epsilon_0) \cap E_f^c} y d\mu \\
&= \int_{\mathcal{B}(x, \epsilon_0) \cap E_f^c} y d\mu + \int_{\mathcal{B}(x, \epsilon_0) \cap E_f^c} y d\mu \\
&= 2 \int_{\mathcal{B}(x, \epsilon_0) \cap E_f^c} y d\mu.
\end{aligned}$$

By inequality (3.4) and equality (3.1), we can conclude that

$$\begin{aligned}
\int_A |f^*(x) - y| d\mu &= 2 \int_{\mathcal{B}(x, \epsilon_0) \cap E_f^c} y d\mu \\
&> 2 \int_{E_f^c} (y - f(x)) d\mu \\
&= \int_A |f(x) - y| d\mu.
\end{aligned}$$

Hence,  $\int_A |f^*(x) - y| d\mu > \int_A |f(x) - y| d\mu$ .

Case ii)  $\int_{E_f^c} f(x) d\mu = 0$  and  $\int_{E_f} (1 - f(x)) d\mu > 0$ .

The function  $H_2$  in Lemma 3.1.2 is continuous on  $(0, \infty)$ .

Consider

$$\begin{aligned}
H_2(0) &= \int_{\mathcal{B}(x,0) \cap E_f} f(z) d\mu - \int_{\mathcal{B}^c(x,0) \cap E_f} (1 - f(z)) d\mu \\
&= 0 - \int_{\mathcal{B}^c(x,0) \cap E_f} (1 - f(z)) d\mu \\
&= - \int_{E_f} (1 - f(z)) d\mu \\
&= \int_{E_f} (f(z) - 1) d\mu \\
&< 0.
\end{aligned}$$

We next show that  $H_2(\infty) > 0$ . Consider

$$\begin{aligned}
H_2(\infty) &= \int_{\mathcal{B}(x,\infty) \cap E_f} f(z) d\mu - \int_{\mathcal{B}^c(x,\infty) \cap E_f} (1 - f(z)) d\mu \\
&= \int_{\mathcal{B}(x,\infty) \cap E_f} f(z) d\mu \\
&= \int_{E_f} f(z) d\mu \\
&> 0.
\end{aligned}$$

Then there is  $\epsilon_0 \in (0, 1)$  such that  $H_2(\epsilon_0) = 0$ . Consider

$$\begin{aligned}
0 &= H_2(\epsilon_0) \\
&= \int_{\mathcal{B}(x,\epsilon_0) \cap E_f} f(z) d\mu - \int_{\mathcal{B}^c(x,\epsilon_0) \cap E_f} (1 - f(z)) d\mu \\
&= \int_{\mathcal{B}(x,\epsilon_0) \cap E_f} f(z) d\mu - \int_{\mathcal{B}^c(x,\epsilon_0) \cap E_f} 1 d\mu + \int_{\mathcal{B}^c(x,\epsilon_0) \cap E_f} f(z) d\mu \\
&= \int_{E_f} f(z) d\mu - \int_{\mathcal{B}^c(x,\epsilon_0) \cap E_f} 1 d\mu.
\end{aligned}$$

Hence,

$$\begin{aligned}
\int_{\mathcal{B}^c(x,\epsilon_0) \cap E_f} 1 d\mu &= \int_{E_f} f(z) d\mu \\
&= y.
\end{aligned} \tag{3.5}$$

We define a new function  $f^* = f \mathbf{1}_{E_f^c} + \mathbf{1}_{\mathcal{B}^c(x,\epsilon_0) \cap E_f}$ . We show that  $f^* \in \mathfrak{D}_y$ .

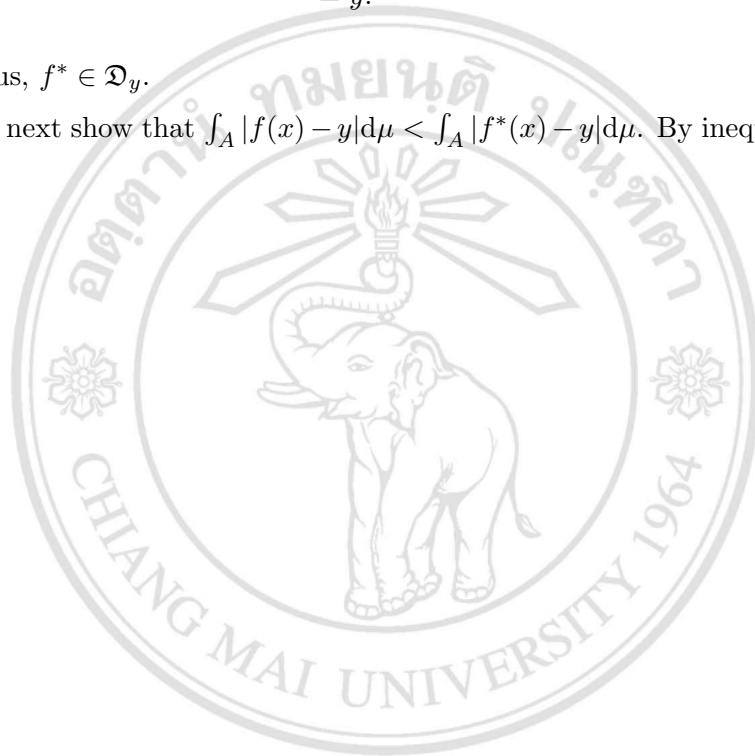


Consider

$$\begin{aligned}
\int_A f^*(x) d\mu &= \int_A f \mathbf{1}_{E_f^c} d\mu + \int_A \mathbf{1}_{\mathcal{B}^c(x, \epsilon_0) \cap E_f} d\mu \\
&= \int_{E_f^c} f(x) d\mu + \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f} 1 d\mu \\
&= 0 + \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f} 1 d\mu \\
&= \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f} 1 d\mu \\
&= y.
\end{aligned}$$

Thus,  $f^* \in \mathfrak{D}_y$ .

We next show that  $\int_A |f(x) - y| d\mu < \int_A |f^*(x) - y| d\mu$ . By inequality (3.2) and



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$\int_{E_f^c} f(x) d\mu = 0$ , we obtain that

$$\begin{aligned}
\int_A |f^*(x) - y| d\mu &= \int_A |f \mathbf{1}_{E_f^c}(x) + \mathbf{1}_{\mathcal{B}^c(x, \epsilon_0) \cap E_f}(x) - y| d\mu \\
&= \int_{E_f} |\mathbf{1}_{\mathcal{B}^c(x, \epsilon_0) \cap E_f}(x) - y| d\mu + \int_{E_f^c} |f(x) - y| d\mu \\
&= \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f} (1 - y) d\mu + \int_{\mathcal{B}(x, \epsilon_0) \cap E_f} y d\mu + \int_{E_f^c} (y - f(x)) d\mu \\
&= \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f} 1 d\mu - \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f} y d\mu + \int_{\mathcal{B}(x, \epsilon_0) \cap E_f} y d\mu \\
&\quad + \int_{E_f^c} y d\mu - \int_{E_f^c} f(x) d\mu \\
&= y - \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f} y d\mu + \int_{\mathcal{B}(x, \epsilon_0) \cap E_f} y d\mu + \int_{E_f^c} y d\mu - 0 \\
&= \int_{E_f^c} y d\mu + \int_{\mathcal{B}(x, \epsilon_0) \cap E_f} y d\mu + \int_{\mathcal{B}(x, \epsilon_0) \cap E_f} y d\mu + \int_{E_f^c} y d\mu \\
&= 2 \left( \int_{E_f^c} y d\mu + \int_{\mathcal{B}(x, \epsilon_0) \cap E_f} y d\mu \right) \\
&> 2 \int_{E_f^c} y d\mu \\
&= 2 \int_{E_f^c} (y - f(x)) d\mu \\
&= \int_{E_f} (f(x) - y) d\mu + \int_{E_f^c} (y - f(x)) d\mu \\
&= \int_{E_f} f(x) d\mu - \int_{E_f} y d\mu + \int_{E_f^c} y d\mu - \int_{E_f^c} f(x) d\mu \\
&= \int_{E_f} |f(x) - y| d\mu + \int_{E_f^c} |f(x) - y| d\mu \\
&= \int_A |f(x) - y| d\mu.
\end{aligned}$$

Thus,  $\int_A |f(x) - y| d\mu < \int_A |f^*(x) - y| d\mu$ . Therefore, the supremum of  $\int_A |f(x) - y| d\mu$  over  $f \in \mathfrak{D}_y$  happens when  $f \in \{0, 1\}$  a.e.

□

**Lemma 3.1.4.** *Let  $Y$  be a continuous random vector with dimension  $n$ . Then*

$$F_{Y|Y}(\vec{y}|\vec{x}) = \begin{cases} 1 & \text{if } x_i < y_i \ ; i = 1, \dots, n \\ 0 & \text{otherwise.} \end{cases}$$

*Proof.* Let  $Y$  be a continuous random vector of dimension  $n$ , we have two cases to consider

Case i)  $\vec{x} < \vec{y}$ .

Then  $V_{F_{Y,Y}}((x-h, x+h] \times (-\infty, y]) = V_{F_Y}((x-h, x+h])$  where  $h$  is sufficiently small. Therefore,

$$\begin{aligned} F_{Y|Y}(y|x) &= \lim_{h \searrow 0} \frac{V_{F_{Y,Y}}((x-h, x+h] \times (-\infty, y])}{V_{F_Y}((x-h, x+h])} \\ &= \lim_{h \searrow 0} \frac{V_{F_Y}((x-h, x+h])}{V_{F_Y}((x-h, x+h])} \\ &= 1. \end{aligned}$$

Case ii)  $x \not< y$ .

If  $x_i > y_i$  for some  $i = 1, \dots, n$ , then  $\lim_{h \searrow 0} V_{F_{Y,Y}}((x-h, x+h] \times (-\infty, y]) = 0$ .

Assume  $x_i \leq y_i$  for all  $i = 1, \dots, n$  and  $x_j = y_j$  for some  $j = 1, \dots, n$ .

Since  $Y$  is a continuous random vector, we get

$$\lim_{h \searrow 0} V_{F_{Y,Y}}((x-h, x+h] \times (-\infty, y]) = 0.$$

□

Recall the definition of  $\varphi$  from Definition 3.1.1,

$$\varphi(Y|X) = \int \int |F_{Y|X}(v|u) - F_Y(v)| dF_X(u) dF_Y(v).$$

**Lemma 3.1.5.** *Let  $Y$  be a continuous random vector. Then*

$$\varphi(Y|Y) = 2 \int (1 - F_Y(y)) F_Y(y) dF_Y(y).$$

*Proof.*

$$\begin{aligned} \varphi(Y|Y) &= \int \int |F_{Y|Y}(y|x) - F_Y(y)| dF_Y(x) dF_Y(y) \\ &= \int \int_{\{x < y\}} (1 - F_Y(y)) dF_Y(x) dF_Y(y) + \int \int_{\{x \not< y\}} F_Y(y) dF_Y(x) dF_Y(y) \\ &= \int \left( (1 - F_Y(y)) \int_{\{x < y\}} dF_Y(x) \right) dF_Y(y) + \int \left( F_Y(y) \int_{\{x \not< y\}} dF_Y(x) \right) dF_Y(y) \\ &= \int (1 - F_Y(y)) F_Y(y) dF_Y(y) + \int (1 - F_Y(y)) F_Y(y) dF_Y(y) \\ &= 2 \int (1 - F_Y(y)) F_Y(y) dF_Y(y). \end{aligned}$$

□

**Lemma 3.1.6.** Let  $X = (X_1, \dots, X_n)$  be a continuous random vector.

Then  $\mathbb{P}(X \in \overline{\mathcal{B}(x, \epsilon)} \setminus \mathcal{B}(x, \epsilon)) = 0$  for all  $x \in \mathbb{R}^n$  and  $\epsilon > 0$ .

*Proof.* Let  $x = (x_1^0, \dots, x_n^0)$ . We set  $A_i = (x_1, \dots, x_i - \epsilon, \dots, x_n)$  and  $B_i = (x_1, \dots, x_i + \epsilon, \dots, x_n)$  where  $x_j \in [x_j^0 - \epsilon, x_j^0 + \epsilon]$  for all  $j \neq i$ . Thus,  $\overline{\mathcal{B}(x, \epsilon)} \setminus \mathcal{B}(x, \epsilon) = \bigcup_{i=1}^n (A_i \cup B_i)$ . Since a random vector  $X$  is continuous, we get  $\mathbb{P}(A_i \cup B_i) \leq \mathbb{P}(\mathbb{R}^{j-1} \times \{x_i^j\} \times \mathbb{R}^{n-j}) = 0$ . Therefore,  $\mathbb{P}(X \in \overline{\mathcal{B}(x, \epsilon)} \setminus \mathcal{B}(x, \epsilon)) = 0$ .  $\square$

**Theorem 3.1.7.** Let  $X$  and  $Y$  be continuous random vectors.

Then  $0 \leq \varphi(Y|X) \leq \varphi(Y|Y)$ .

*Proof.* It is easy to see that  $0 \leq \varphi(Y|X)$ . Next, we show that  $\varphi(Y|X) \leq \varphi(Y|Y)$ .

By Lemma 3.1.3, we obtain that

$$2 \int (1 - F_Y(y)) F_Y(y) dF_Y(y) \geq \int \int |F_{Y|X}(y|x) - F_Y(y)| dF_X(x) dF_Y(y).$$

By Lemma 3.1.5, we have

$$\begin{aligned} \varphi(Y|Y) &= 2 \int (1 - F_Y(y)) F_Y(y) dF_Y(y) \\ &\geq \int |F_{Y|X}(y|x) - F_Y(y)| dF_X(x) dF_Y(y) \\ &= \varphi(Y|X). \end{aligned}$$

$\square$

**Theorem 3.1.8.** Let  $X$  and  $Y$  be continuous random vectors. Then  $\varphi(Y|X) = 0$  if and only if  $X$  and  $Y$  are independent.

*Proof.* Assume that  $\varphi(Y|X) = 0$ . Then  $\int \int |F_{Y|X}(v|u) - F_Y(v)| dF_X(u) dF_Y(v) = 0$ .

Thus,  $F_{Y|X}(v|u) - F_Y(v) = 0$ . Hence,  $F_{Y|X}(v|u) = F_Y(v)$ .

Therefore,  $X$  and  $Y$  are independent.

Conversely, assume  $X$  and  $Y$  are independent.

We get  $F_{Y|X}(v|u) = F_Y(v)$ . Thus,

$$\begin{aligned} \varphi(Y|X) &= \int |F_{Y|X}(v|u) - F_Y(v)| dF_X(u) dF_Y(v) \\ &= \int |F_Y(v) - F_Y(v)| dF_X(u) dF_Y(v) \\ &= 0. \end{aligned}$$

$\square$

**Theorem 3.1.9.** *Let  $X$  and  $Y$  be continuous random vectors.*

*Then  $\varphi(Y|X) = \varphi(Y|Y)$  if and only if  $Y$  is a function of  $X$ .*

*Proof.* Assume that  $\varphi(Y|X) = \varphi(Y|Y)$ .

Then  $\int \int |F_{Y|X}(y|x) - F_Y(y)| dF_X(x) dF_Y(y) = \int 2(1 - F_Y(y)) F_Y(y) dF_Y(y)$ .

Since  $\int |F_{Y|X}(y|x) - F_Y(y)| dF_X \leq 2(1 - F_Y(y)) F_Y(y)$ , we can conclude that

$$\int |F_{Y|X}(y|x) - F_Y(y)| dF_X = 2(1 - F_Y(y)) F_Y(y).$$

By Lemma 3.1.3 and Lemma 2.2.4, we have  $F_{Y|X} \in \{0, 1\}$  a.e.

Since  $F_{Y|X}(\cdot|x)$  is a distribution function, we obtain that the set  $\{y | F_{Y|X}(y|x) = 1\}$  is closed under the pointwise minimum and closed under Euclidean topology.

Therefore,  $\{y | F_{Y|X}(y|x) = 1\}$  has a minimum, we say  $f(x)$ .

Then  $F_{Y|X}(\cdot|x) = \mathbf{1}_{[f(x), \infty]}$ . Thus,  $\mathbb{P}(Y = f(X)) = 1$ .

Hence,  $Y$  is a function of  $X$ . □

In fact,  $\varphi(Y|X)$  is not generally less than one so, we normalize it into the form

$$\frac{\varphi(Y|X)}{\varphi(Y|Y)} = \frac{\int |F_{Y|X}(y|x) - F_Y(y)| dF_X(x) dF_Y(y)}{2 \int F_Y(y)(1 - F_Y(y)) dF_Y(y)}.$$

For any random vectors  $X$  and  $Y$  with the joint distribution function  $F_{X,Y}$ , we denote  $\varphi(F_{X,Y}|X) = \varphi(Y|X)$ .

Let  $F$  and  $G$  be joint distribution functions with the same marginals, then the convex combination  $tF + (1 - t)G$  also has the same marginals as  $F$  and  $G$ .

**Theorem 3.1.10.** *Let  $F$  and  $G$  be joint distribution functions with the same marginals. Let  $(X_1, Y_1), (X_2, Y_2)$  and  $(X_t, Y_t)$  have the distribution functions  $F, G$  and  $tF + (1 - t)G$ , respectively. Then  $\varphi(Y_t|X_t) \leq t\varphi(Y_1|X_1) + (1 - t)\varphi(Y_2|X_2)$ .*

*Proof.* Let  $H = tF + (1 - t)G$ .

Then  $H_{Y_t|X_t} = tF_{Y_1|X_1} + (1 - t)G_{Y_2|X_2}$  and  $H_{Y_t} = tF_{Y_1} + (1 - t)G_{Y_2}$ . Thus,

$$\begin{aligned} \varphi(H|X_t) &= \int |H_{Y_t|X_t}(y|x) - H_{Y_t}(y)| dF_X(x) dF_Y(y) \\ &= \int |tF_{Y_1|X_1}(y|x) + (1 - t)G_{Y_2|X_2}(y|x) - tF_{Y_1}(y) - (1 - t)G_{Y_2}(y)| dF_X(x) dF_Y(y) \\ &= \int |t(F_{Y_1|X_1}(y|x) - F_{Y_1}(y)) + (1 - t)(G_{Y_2|X_2}(y|x) - G_{Y_2}(y))| dF_X(x) dF_Y(y) \\ &\leq t \int |F_{Y_1|X_1}(y|x) - F_{Y_1}(y)| dF_X(x) dF_Y(y) \\ &\quad + (1 - t) \int |G_{Y_2|X_2}(y|x) - G_{Y_2}(y)| dF_X(x) dF_Y(y) \\ &= t\varphi(F|X_1) + (1 - t)\varphi(G|X_2). \end{aligned}$$

□

By Theorem 3.1.10, we get  $\varphi(tF_{X,Y} + (1-t)G_{X,Y}|X) \leq \max\{\varphi(F_{X,Y}|X), \varphi(G_{X,Y}|X)\}$  for all joint distributions  $F_{X,Y}$  and  $G_{X,Y}$  with the same marginals  $F_X$  and  $F_Y$ .

**Theorem 3.1.11.** *Let  $X$  and  $Y$  be random vectors.*

*Then  $\varphi(Y|f(X)) \leq \varphi(Y|X)$  for all measurable functions  $f$ .*

*Proof.* Let  $f(X) = Z$ . We get

$$\begin{aligned} \varphi(Y|X) &= \int \int |F_{Y|X}(y|x) - F_Y(y)| dF_X(x) dF_Y(y) \\ &= \int \int \int |F_{Y|X}(y|x) - F_Y(y)| dF_{X|Z}(x|z) dF_Z(z) dF_Y(y) \\ &\geq \int \int \left| \int F_{Y|X}(y|x) dF_{X|Z}(x|z) - F_Y(y) \right| dF_Z(z) dF_Y(y) \\ &= \int \int |F_{Y|Z}(y|z) - F_Y(y)| dF_Z(z) dF_Y(y) \\ &= \varphi(Y|Z). \end{aligned}$$

Hence,  $\varphi(Y|f(X)) \leq \varphi(Y|X)$  for all measurable functions  $f$ . □

**Theorem 3.1.12.** *Let  $X, Y$  and  $Z$  be random vectors. Then  $\varphi(Y, Y, Z|X) = \varphi(Y, Z|X)$ .*

*Proof.* For any random vectors  $X, Y$  and  $Z$  we have

$$\begin{aligned} \varphi(Y, Y, Z|X) &= \int \int |F_{Y,Y,Z|X}(y, w, z|x) - F_{Y,Y,Z}(y, w, z)| dF_{Y,Y,Z}(y, w, z) dF_X(x) \\ &= \int \int |F_{Y,Z|X}(y, z|x) - F_{Y,Z}(y, z)| dF_{Y,Z}(y, z) dF_X(x) \\ &= \varphi(Y, Z|X). \end{aligned}$$

Therefore,  $\varphi(Y, Y, Z|X) = \varphi(Y, Z|X)$ . □

### 3.2 Measure of Complete Dependence Based on Linkages

**Definition 3.2.1.** Let  $X$  and  $Y$  be absolutely continuous random vectors of dimensions  $m$  and  $n$ , respectively and  $p \in [1, \infty)$ .

We define the measure  $\zeta_p$  of complete dependence by

$$\zeta_p(Y|X) = \left[ \int_{[0,1]^m} \int_{[0,1]^n} \left| \frac{\partial}{\partial u} C_{X,Y}(\vec{u}, \vec{v}) - \Pi(\vec{v}) \right|^p d\vec{u} d\vec{v} \right]^{\frac{1}{p}}$$

where  $C_{X,Y}$  is the linkage associated with  $(X, Y)$ .

**Theorem 3.2.1.** *Let  $X$  and  $Y$  be two absolutely continuous random vectors and  $f$  be a measurable function such that  $f(X)$  has an absolutely continuous distribution function. Then  $\zeta_p(Y|f(X)) \leq \zeta_p(Y|X)$ .*

*Proof.* Since  $\zeta_p(Y|X) = \zeta_p(\Psi_{F_Y}(Y)|\Psi_{F_X}(X))$ , we may as well assume that  $X$  and  $Y$  are uniformly distributed. Similarly, the fact that  $\zeta_p(Y|f(X)) = \zeta_p(Y|\Psi_{F_{f(X)}}(f(X)))$  allows us to assume that  $Z = f(X)$  has uniform distribution also. Now,

$$\begin{aligned} \frac{\partial}{\partial \vec{z}} C_{Z,Y}(\vec{z}, \vec{y}) &= F_{Y|Z}(\vec{y}|\vec{z}) \\ &= \int F_{Y|X}(\vec{y}|\vec{x}) dF_{X|Z}(\vec{x}|\vec{z}) \\ &= \int \frac{\partial}{\partial \vec{x}} C_{X,Y}(\vec{x}, \vec{y}) dF_{X|Z}(\vec{x}|\vec{z}). \end{aligned}$$

By Jensen's inequality, we have

$$\begin{aligned} \int \int \left| \frac{\partial}{\partial \vec{z}} C_{Z,Y}(\vec{z}, \vec{y}) - \Pi(\vec{y}) \right|^p d\vec{z} d\vec{y} &= \int \int \left| \int \frac{\partial}{\partial \vec{x}} C_{X,Y}(\vec{x}, \vec{y}) dF_{X|Z}(\vec{x}|\vec{z}) - \Pi(\vec{y}) \right|^p d\vec{z} d\vec{y} \\ &= \int \int \left| \int \left( \frac{\partial}{\partial \vec{x}} C_{X,Y}(\vec{x}, \vec{y}) - \Pi(\vec{y}) \right) dF_{X|Z}(\vec{x}|\vec{z}) \right|^p d\vec{z} d\vec{y} \\ &\leq \int \int \int \left| \frac{\partial}{\partial \vec{x}} C_{X,Y}(\vec{x}, \vec{y}) - \Pi(\vec{y}) \right|^p dF_{X|Z}(\vec{x}|\vec{z}) d\vec{z} d\vec{y} \\ &= \int \int \left| \frac{\partial}{\partial \vec{x}} C_{X,Y}(\vec{x}, \vec{y}) - \Pi(\vec{y}) \right|^p dF_X(\vec{x}) d\vec{y}, \end{aligned}$$

that is,  $\zeta_p(Y|Z) \leq \zeta_p(Y|X)$ . □

**Lemma 3.2.2.** *Let  $X$  and  $Y$  be two absolutely continuous random vectors.*

*Then the three following properties are equivalent:*

- i)  $\zeta_p(Y|X) = 0$ ,
- ii)  $\frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v}) = \Pi(\vec{v})$ , and
- iii)  $C_{X,Y}(\vec{u}, \vec{v}) = \Pi(\vec{u})\Pi(\vec{v})$ .

*Proof.* i)  $\Rightarrow$  ii). Assume that  $\zeta_p(Y|X) = 0$ .

Then  $\left[ \int_{[0,1]^m} \int_{[0,1]^n} \left| \frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v}) - \Pi(\vec{v}) \right|^p d\vec{u} d\vec{v} \right]^{\frac{1}{p}} = 0$ .

Thus,  $\int_{[0,1]^m} \int_{[0,1]^n} \left| \frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v}) - \Pi(\vec{v}) \right|^p d\vec{u} d\vec{v} = 0$ .

Hence,  $\left| \frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v}) - \Pi(\vec{v}) \right| = 0$ .

Therefore,  $\frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v}) = \Pi(\vec{v})$ .

ii)  $\Rightarrow$  iii). Assume that  $\frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v}) = \Pi(\vec{v})$ . Thus,

$$\begin{aligned} C_{X,Y}(\vec{a}, \vec{b}) &= \int_{[\vec{0}, \vec{a}]} \frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{b}) d\vec{u} \\ &= \int_{[\vec{0}, \vec{a}]} \Pi(\vec{b}) d\vec{u} \\ &= \Pi(\vec{a})\Pi(\vec{b}). \end{aligned}$$

iii)  $\Rightarrow$  i). Assume that  $C_{X,Y}(\vec{u}, \vec{v}) = \Pi(\vec{u})\Pi(\vec{v})$ . Then

$$\begin{aligned} \zeta_p(Y|X) &= \left[ \int_{[0,1]^m} \int_{[0,1]^n} \left| \frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v}) - \Pi(\vec{v}) \right|^p d\vec{u} d\vec{v} \right]^{\frac{1}{p}} \\ &= \left[ \int_{[0,1]^m} \int_{[0,1]^n} \left| \frac{\partial}{\partial \vec{u}} \Pi(\vec{u})\Pi(\vec{v}) - \Pi(\vec{v}) \right|^p d\vec{u} d\vec{v} \right]^{\frac{1}{p}} \\ &= \left[ \int_{[0,1]^m} \int_{[0,1]^n} |\Pi(\vec{v}) - \Pi(\vec{v})|^p d\vec{u} d\vec{v} \right]^{\frac{1}{p}} \\ &= 0. \end{aligned}$$

□

**Lemma 3.2.3.** Let  $X$  and  $Y$  be two absolutely continuous random vectors. Then  $X$  and  $Y$  are independent if and only if  $\Psi_{F_X}(X)$  and  $\Psi_{F_Y}(Y)$  are independent.

*Proof.* Since  $X$  and  $Y$  are absolutely continuous random vectors, we have  $\Psi_{F_X}$  and  $\Psi_{F_Y}$  are invertible functions. Thus,

$$\begin{aligned} \mathbb{P}(\Psi_{F_X}(X) \in A, \Psi_{F_Y}(Y) \in B) &= \mathbb{P}(X \in \Psi_{F_X}^{-1}(A), Y \in \Psi_{F_Y}^{-1}(B)) \\ &= \mathbb{P}(X \in \Psi_{F_X}^{-1}(A))\mathbb{P}(Y \in \Psi_{F_Y}^{-1}(B)) \\ &= \mathbb{P}(\Psi_{F_X}(X) \in A)\mathbb{P}(\Psi_{F_Y}(Y) \in B). \end{aligned}$$

Therefore,  $\Psi_{F_X}(X)$  and  $\Psi_{F_Y}(Y)$  are independent.

Conversely, suppose that  $\Psi_{F_X}(X)$  and  $\Psi_{F_Y}(Y)$  are independent. Consider

$$\begin{aligned} \mathbb{P}(X \in A, Y \in B) &= \mathbb{P}(X \in \Psi_{F_X}^{-1}(\Psi_{F_X}(A)), Y \in \Psi_{F_Y}^{-1}(\Psi_{F_Y}(B))) \\ &= \mathbb{P}(\Psi_{F_X}(X) \in \Psi_{F_X}(A), \Psi_{F_Y}(Y) \in \Psi_{F_Y}(B)) \\ &= \mathbb{P}(\Psi_{F_X}(X) \in \Psi_{F_X}(A))\mathbb{P}(\Psi_{F_Y}(Y) \in \Psi_{F_Y}(B)) \\ &= \mathbb{P}(X \in A)\mathbb{P}(Y \in B). \end{aligned}$$

Hence,  $X$  and  $Y$  are independent.

□



**Theorem 3.2.4.** *Let  $X$  and  $Y$  be two absolutely continuous random vectors.*

*Then  $\zeta_p(Y|X) = 0$  if and only if  $X$  and  $Y$  are independent.*

*Proof.* Assume that  $\zeta_p(Y|X) = 0$ . By Lemma 3.2.2, we have  $C_{X,Y}(\vec{u}, \vec{v}) = \Pi(\vec{u})\Pi(\vec{v})$ . Since  $C_{X,Y}$  is the joint distribution of  $\Psi_{F_X}(X)$  and  $\Psi_{F_Y}(Y)$ , we obtain that  $\Psi_{F_X}(X)$  and  $\Psi_{F_Y}(Y)$  are independent. By Lemma 3.2.3, we get  $X$  and  $Y$  are independent.

Conversely, let  $X$  and  $Y$  be independent. By Lemma 3.2.2, we get  $\Psi_{F_X}(X)$  and  $\Psi_{F_Y}(Y)$  are independent. Thus,  $C_{X,Y} = \Pi(\vec{u})\Pi(\vec{v})$ . By Lemma 3.2.3, we can conclude that  $\zeta_p(Y|X) = 0$ .  $\square$

**Lemma 3.2.5.** *Let  $A$  be a metric space and  $\mu$  be a Borel probability measure on  $A$  such that  $\mu(\overline{\mathcal{B}(x, \epsilon)} \setminus \mathcal{B}(x, \epsilon)) = 0$  for all ball  $\mathcal{B}(x, \epsilon)$  with centered in  $x \in A$  and of radius  $\epsilon > 0$ . Let  $y \in (0, 1)$  and  $\mathfrak{D}_y = \{f : A \rightarrow [0, 1] \mid f \text{ is measurable and } \int_A f(x) d\mu = y\}$ . Then  $f$  maximizes the function  $\hat{f} \mapsto \int |\hat{f}(x) - y|^p d\mu$  on  $\mathfrak{D}_y$  if and only if  $f$  is an indicator function. Moreover,  $\max_{\{f \in \mathfrak{D}_y\}} \int_A |f(x) - y|^p d\mu = y^p(1 - y) + y(1 - y)^p$ .*

*Proof.* By symmetry, we may assume  $0 < y \leq \frac{1}{2}$ . By following the proof of Lemma 3.1.3, we have  $\mathfrak{D}_y$  contains an indicator function.

Let  $B \subseteq A$  and  $\mathbf{1}_B \in \mathfrak{D}_y$  be an indicator function of  $B$ . We shall show that

$$\int_A |\mathbf{1}_B(x) - y|^p d\mu = (1 - y)^p y + y^p(1 - y).$$

Since  $\mathbf{1}_B \in \mathfrak{D}_y$ , we can conclude that  $\mu(B) = \int \mathbf{1}_B d\mu = y$ . Consider

$$\begin{aligned} \int_A |\mathbf{1}_B(x) - y|^p d\mu &= \int_B (1 - y)^p d\mu + \int_{B^c} y^p d\mu \\ &= (1 - y)^p \mu(B) + y^p \mu(B^c) \\ &= (1 - y)^p y + y^p(1 - y). \end{aligned}$$

Therefore,  $\int_A |\mathbf{1}_B(x) - y|^p d\mu = (1 - y)^p y + y^p(1 - y)$ .

Let  $f \in \mathfrak{D}_y$  and  $E_f = \{x \in A \mid f(x) > y\}$ .

Assume that  $f$  is not an indicator function.

We will construct another function  $f^* \in \mathfrak{D}_y$  such that  $\int_A |f(x) - y|^p d\mu < \int_A |f^*(x) - y|^p d\mu$ .

Hence,  $f$  can not maximize  $\hat{f} \mapsto \int |\hat{f}(x) - y|^p d\mu$  on  $\mathfrak{D}_y$ .

Since  $f \in \mathfrak{D}_y$ ,  $\int_{E_f^c} (y - f(x)) d\mu = \int_{E_f} (f(x) - y) d\mu$ .

If  $\int_{E_f^c} f(x) d\mu = 0$ , then  $\int_{E_f} f(x) d\mu = y$  and hence,  $f = y$  a.e. which immediately implies  $f$  is not the maximizer.

Thus, we may assume  $\int_{E_f^c} f(x) d\mu > 0$ .

For any  $\epsilon > 0$ , define  $H(\epsilon) = \int_{\mathcal{B}(x,\epsilon) \cap E_f^c} f(z) d\mu - \int_{\mathcal{B}^c(x,\epsilon) \cap E_f^c} (1 - f(z)) d\mu$ . Since  $0 > y - 1 \geq -(1 - f(x))$ , we get  $H(0) = - \int_{E_f^c} (1 - f(z)) d\mu < 0$ .

Clearly,  $\lim_{\epsilon \rightarrow \infty} H(\epsilon) = \int_{E_f^c} f(z) d\mu > 0$ .

By the assumption of  $\mu$ , the function  $H$  is continuous.

Thus, there exists  $\epsilon_0 \in (0, \infty)$  such that  $H(\epsilon_0) = 0$ .

Define a new function  $f^*$  by letting  $f^* = f \mathbf{1}_{E_f} + \mathbf{1}_{E_f^c \cap \mathcal{B}^c(x, \epsilon_0)}$ .

First, we show that  $f^* \in \mathfrak{D}_y$ . Since  $H(\epsilon_0) = 0$ , we have

$$\begin{aligned} 0 &= \int_{\mathcal{B}(x, \epsilon_0) \cap E_f^c} f(z) d\mu - \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c} (1 - f(z)) d\mu \\ &= \int_{\mathcal{B}(x, \epsilon_0) \cap E_f^c} f(z) d\mu - \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c} 1 d\mu + \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c} f(z) d\mu. \end{aligned}$$

Therefore,

$$\int_{E_f^c} f(z) d\mu = \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c} 1 d\mu. \quad (3.6)$$

Then

$$\begin{aligned} \int_A f^*(x) d\mu &= \int_A f \mathbf{1}_{E_f}(x) d\mu + \int_A \mathbf{1}_{E_f^c \cap \mathcal{B}^c(x, \epsilon_0)}(x) d\mu \\ &= \int_{E_f} f(x) d\mu + \int_{E_f^c \cap \mathcal{B}^c(x, \epsilon_0)} 1 d\mu \\ &= \int_{E_f} f(x) d\mu + \int_{E_f^c} f(z) d\mu \\ &= \int_A f(x) d\mu \\ &= y. \end{aligned}$$

Hence,  $f^* \in \mathfrak{D}_y$ . Next, we show that  $\int_A |f(x) - y|^p d\mu < \int_A |f^*(x) - y|^p d\mu$ . Consider

$$\begin{aligned} \int_A |f^*(x) - y|^p d\mu &= \int_{E_f} |f^*(x) - y|^p d\mu + \int_{E_f^c} |f^*(x) - y|^p d\mu \\ &= \int_{E_f} |f \mathbf{1}_{E_f}(x) + \mathbf{1}_{E_f^c \cap \mathcal{B}^c(x, \epsilon_0)}(x) - y|^p d\mu \\ &\quad + \int_{E_f^c} |f \mathbf{1}_{E_f}(x) + \mathbf{1}_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c}(x) - y|^p d\mu \\ &= \int_{E_f} |f(x) - y|^p d\mu + \int_{E_f^c} |\mathbf{1}_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c}(x) - y|^p d\mu \\ &= \int_{E_f} |f(x) - y|^p d\mu + \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c} (1 - y)^p d\mu + \int_{\mathcal{B}(x, \epsilon_0) \cap E_f^c} y^p d\mu. \end{aligned}$$

Clearly,  $|f(x) - y| \leq y \leq 1 - y$  for all  $x \in E^c$ . Thus,

$$\int_{E_f^c} |f(x) - y|^p d\mu \leq \int_{\mathcal{B}^c(x, \epsilon_0) \cap E_f^c} (1 - y)^p d\mu + \int_{\mathcal{B}(x, \epsilon_0) \cap E_f^c} y^p d\mu$$

and both sides are equal only when  $y = \frac{1}{2}$  and in this case  $|f(x) - y| = y$  on  $E_f^c$ , that is,  $f = 0$  on  $E_f^c$ . This contradicts the fact  $\int_{E_f^c} f(x) d\mu > 0$ . Thus,

$$\int_{E_f^c} |f(x) - y|^p d\mu < \int_{B^c(x, \epsilon_0) \cap E_f^c} (1 - y)^p d\mu + \int_{B(x, \epsilon_0) \cap E_f^c} y^p d\mu.$$

Finally, we show that

$$\int_A |f(x) - y|^p d\mu = (1 - y)^p y + y^p (1 - y)$$

whenever  $f \in \mathfrak{D}_y$  is an indicator function. Consider

$$\begin{aligned} \int_A |f(x) - y|^p d\mu &= \int_{E_f} |f(x) - y|^p d\mu + \int_{E_f^c} |f(x) - y|^p d\mu \\ &= \int_{E_f} (1 - y)^p d\mu + \int_{E_f^c} y^p d\mu \\ &= (1 - y)^p \mu(E_f) + y^p \mu(E_f^c) \\ &= (1 - y)^p y + y^p (1 - y). \end{aligned}$$

□

**Corollary 3.2.6.** *For any absolutely continuous random vector  $Y$ ,*

$$\zeta_p(Y|Y) = \left[ \int_{[0,1]^n} (\Pi(\vec{v}) (1 - \Pi(\vec{v}))^p + (\Pi(\vec{v}))^p (1 - \Pi(\vec{v}))) d\vec{v} \right]^{\frac{1}{p}}.$$

*Proof.* Since  $\frac{\partial}{\partial \vec{u}} C_{Y,Y}(\vec{u}, \vec{v}) = C_{Y|Y}(\vec{v}|\vec{u}) = \mathbf{1}_{\{(\vec{u}, \vec{v}) | \vec{v} \leq \vec{u}\}}$  is an indicator function,

$$\int_{[0,1]^m} \left| \frac{\partial}{\partial \vec{u}} C_{Y,Y}(\vec{u}, \vec{v}) - \Pi(\vec{v}) \right|^p d\vec{u} = \Pi(\vec{v}) (1 - \Pi(\vec{v}))^p + (\Pi(\vec{v}))^p (1 - \Pi(\vec{v})),$$

by Lemma 3.2.5. Thus,

$$\zeta_p(Y|Y) = \left[ \int_{[0,1]^n} (\Pi(\vec{v}) (1 - \Pi(\vec{v}))^p + (\Pi(\vec{v}))^p (1 - \Pi(\vec{v}))) d\vec{v} \right]^{\frac{1}{p}}.$$

□

**Lemma 3.2.7.** *Let  $Y$  be an absolutely continuous  $m$ -dimensional random vector and  $p$  be a positive integer. Then*

$$\zeta_p(Y|Y)^p = \sum_{k=0}^p \binom{p}{k} (-1)^k \frac{1}{(k+2)^m} + \frac{1}{(p+1)^m} - \frac{1}{(p+2)^m}.$$

*Proof.* Consider

$$\begin{aligned}
\zeta_p(Y|Y)^p &= \int_0^1 \dots \int_0^1 (\Pi(y)(1 - \Pi(y))^p + \Pi(y)^p(1 - \Pi(y))) dy_1 \dots dy_m \\
&= \int_0^1 \dots \int_0^1 \left( \Pi(y) \sum_{k=0}^p \binom{p}{k} (-\Pi(y))^k + \Pi(y)^p - \Pi(y)^{p+1} \right) dy_1 \dots dy_m \\
&= \int_0^1 \dots \int_0^1 \left( \sum_{k=0}^p \binom{p}{k} (-1)^k \Pi(y)^{k+1} + \Pi(y)^p - \Pi(y)^{p+1} \right) dy_1 \dots dy_m \\
&= \sum_{k=0}^p \binom{p}{k} (-1)^k \frac{1}{(k+2)^m} + \frac{1}{(p+1)^m} - \frac{1}{(p+2)^m}.
\end{aligned}$$

□

**Remark 3.2.8.** Let  $Y$  be a random vector with dimension  $m$ , by Lemma 3.2.7, we get

$$\begin{aligned}
\zeta_1(Y|Y) &= \sum_{k=0}^1 \binom{1}{k} (-1)^k \frac{1}{(k+2)^m} + \frac{1}{(1+1)^m} - \frac{1}{(1+2)^m} \\
&= \binom{1}{0} \frac{1}{2^m} - \binom{1}{1} \frac{1}{3^m} + \frac{1}{2^m} - \frac{1}{3^m} \\
&= \frac{1}{2^m} - \frac{1}{3^m} + \frac{1}{2^m} - \frac{1}{3^m} \\
&= \frac{2}{2^m} - \frac{2}{3^m}
\end{aligned}$$

and

$$\begin{aligned}
\zeta_2^2(Y|Y) &= \sum_{k=0}^2 \binom{2}{k} (-1)^k \frac{1}{(k+2)^m} + \frac{1}{(2+1)^m} - \frac{1}{(2+2)^m} \\
&= \binom{2}{0} \frac{1}{2^m} - \binom{2}{1} \frac{1}{3^m} + \binom{2}{2} \frac{1}{4^m} + \frac{1}{3^m} - \frac{1}{4^m} \\
&= \frac{1}{2^m} - \frac{2}{3^m} + \frac{1}{4^m} + \frac{1}{3^m} - \frac{1}{4^m} \\
&= \frac{1}{2^m} - \frac{1}{3^m}.
\end{aligned}$$

Therefore,  $\zeta_1(Y|Y) = \frac{2}{2^m} - \frac{2}{3^m}$  and  $\zeta_2(Y|Y) = \sqrt{\frac{2}{2^m} - \frac{2}{3^m}}$ .

**Theorem 3.2.9.** Let  $X$  and  $Y$  be two absolutely continuous random vectors.

Then  $0 \leq \zeta_p(Y|X) \leq \zeta_p(Y|Y)$ .

*Proof.* Clearly,  $0 \leq \zeta_p(Y|X)$ . Next, we show that  $\zeta_p(Y|X) \leq \zeta_p(Y|Y)$ .

By Lemma 3.2.5,

$$\int \left| \frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v}) - \Pi(\vec{v}) \right|^p d\vec{u} \leq (\Pi(\vec{v}))^p (1 - \Pi(\vec{v})) + \Pi(\vec{v})(1 - \Pi(\vec{v}))^p.$$

Therefore,

$$\begin{aligned}\zeta_p(Y|X) &= \left( \int \int \left| \frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v}) - \Pi(\vec{v}) \right|^p d\vec{u} d\vec{v} \right)^{\frac{1}{p}} \\ &\leq \left( \int [(\Pi(\vec{v}))^p (1 - \Pi(\vec{v})) + \Pi(\vec{v})(1 - \Pi(\vec{v}))^p] d\vec{v} \right)^{\frac{1}{p}} \\ &= \zeta_p(Y|Y).\end{aligned}$$

□

**Theorem 3.2.10.** *Let  $X$  and  $Y$  be two absolutely continuous random vectors.*

*The three following properties are equivalent:*

- i)  $Y$  is a measurable function of  $X$ ,*
- ii)  $\Psi_{F_Y}(Y)$  is a measurable function of  $\Psi_{F_X}(X)$ , and*
- iii)  $\zeta_p(Y|X) = \zeta_p(Y|Y)$ .*

*Proof.* First, we show that *i)* and *ii)* are equivalent. Assume  $Y$  is a measurable function of  $X$ , that is,  $Y = f(X)$  for some measurable function  $f$ .

Thus,

$$\Psi_{F_Y}(Y) = \Psi_{F_{f(X)}}(f(X)) = \Psi_{F_{f(X)}} f(\Psi_{F_{f(X)}}^{-1}(\Psi_{F_{f(X)}}(X))),$$

that is,  $\Psi_{F_Y}(Y)$  is a measurable function of  $\Psi_{F_X}(X)$ .

Conversely, assume  $\Psi_{F_Y}(Y)$  is a measurable function of  $\Psi_{F_X}(X)$ .

Then  $\Psi_{F_Y}(Y) = g(\Psi_{F_X}(X))$  for some measurable function  $g$ . So,  $Y = \Psi_{F_Y}^{-1}(g(\Psi_{F_X}(X)))$ .

Finally, we show that *i)* and *iii)* are equivalent.

Assume  $Y$  is a measurable function of  $X$ , that is,  $Y = f(X)$  for some measurable function  $f$ . The fact that  $\zeta_p(Y|X) = \zeta_p(Y|Y)$  immediately follows from Theorem 3.2.1 and Theorem 3.2.9.

Conversely, assume  $\zeta_p(Y|X) = \zeta_p(Y|Y)$ .

$$\text{Then } \int \left| \frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v}) - \Pi(\vec{v}) \right|^p d\vec{u} d\vec{v} = \int [(\Pi(\vec{v}))^p (1 - \Pi(\vec{v})) + \Pi(\vec{v})(1 - \Pi(\vec{v}))^p] d\vec{v}.$$

$$\text{Thus, } \int \left| \frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v}) - \Pi(\vec{v}) \right|^p d\vec{u} = (\Pi(\vec{v}))^p (1 - \Pi(\vec{v})) + \Pi(\vec{v})(1 - \Pi(\vec{v}))^p.$$

By Lemma 3.2.5,  $\frac{\partial}{\partial \vec{u}} C_{X,Y}$  is an indicator function.

Since  $\frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \cdot) = C_{Y|X}(\cdot|\vec{u})$  is a distribution function, the set  $\left\{ \vec{v} \mid \frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v}) = 1 \right\}$  is closed under the pointwise minimum and closed under Euclidean topology.

Therefore,  $\left\{ \vec{v} \mid \frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v}) = 1 \right\}$  has a minimum, say  $f(\vec{u})$ .

Then  $\frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v}) = \mathbf{1}_{[f(\vec{u}), \infty)}(\vec{v})$ .

Thus,  $\mathbb{P}(Y = f(X) = 1)$ , that is,  $Y$  is a measurable function of  $X$ .  $\square$

**Theorem 3.2.11.** *Let  $X$ ,  $Y$ , and  $Z$  be absolutely continuous random vectors in which  $Z$  has the dimension  $k$  and  $(X, Y)$  and  $Z$  are independent,*

$$\zeta_p(Y, Z|X) = \left( \frac{1}{p+1} \right)^{\frac{k}{p}} \zeta_p(Y|X).$$

In particular,  $\zeta_p(Y, Z|X) < \zeta_p(Y|X)$ .

*Proof.* Since  $(X, Y)$  and  $Z$  are independent, we have  $C_{X,(Y,Z)}(\vec{u}, (\vec{v}, \vec{w})) = C_{X,Y}(\vec{u}, \vec{v})\Pi(\vec{w})$ .

Thus,

$$\begin{aligned} (\zeta_p(Y, Z|X))^p &= \int \int \int \left| \frac{\partial}{\partial \vec{u}} C_{X,(Y,Z)}(\vec{u}, (\vec{v}, \vec{w})) - \Pi(\vec{v})\Pi(\vec{w}) \right|^p d\vec{u} d\vec{v} d\vec{w} \\ &= \int \int \int \left| \frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v})\Pi(\vec{w}) - \Pi(\vec{v})\Pi(\vec{w}) \right|^p d\vec{u} d\vec{v} d\vec{w} \\ &= \int (\Pi(\vec{w}))^p d\vec{w} \int \int \left| \frac{\partial}{\partial \vec{u}} C_{X,Y}(\vec{u}, \vec{v}) - \Pi(\vec{v}) \right|^p d\vec{u} d\vec{v} \\ &= \left( \frac{1}{p+1} \right)^k \zeta_p(Y|X)^p. \end{aligned}$$

Therefore,  $\zeta_p(Y, Z|X) = \left( \frac{1}{p+1} \right)^{\frac{k}{p}} \zeta_p(Y|X)$ .  $\square$

**Corollary 3.2.12.** *Let  $X, Y$  and  $Z$  be absolutely continuous random vectors such that  $(X, Y)$  and  $Z$  are independent. Then  $\frac{\zeta_k(Y, Z|X)}{\zeta_k(Y, Z|Y, Z)} < \frac{\zeta_k(Y|X)}{\zeta_k(Y|Y)}$ .*

*Proof.* Assume that  $Y$  has the dimension  $m$  and  $Z$  has the dimension  $n$ . Consider

$$\begin{aligned} \zeta_p^k(Y, Z|Y, Z) &= \int_{[0,1]^m} \int_{[0,1]^n} \left[ \Pi(\vec{u}, \vec{v})(1 - \Pi(\vec{u}, \vec{v}))^k + \Pi^k(\vec{u}, \vec{v})(1 - \Pi(\vec{u}, \vec{v})) \right] d\vec{v} d\vec{u} \\ &= \int_{[0,1]^m} \int_{[0,1]^n} \left[ \Pi(\vec{u})\Pi(\vec{v})(1 - \Pi(\vec{u})\Pi(\vec{v}))^k + \Pi^k(\vec{u})\Pi^k(\vec{v})(1 - \Pi(\vec{u})\Pi(\vec{v})) \right] d\vec{v} d\vec{u} \\ &> \int_{[0,1]^m} \int_{[0,1]^n} \left[ \Pi(\vec{u})\Pi^k(\vec{v})(1 - \Pi(\vec{u}))^k + \Pi^k(\vec{u})\Pi^k(\vec{v})(1 - \Pi(\vec{u})) \right] d\vec{v} d\vec{u} \\ &= \frac{1}{(p+1)^n} \int_{[0,1]^m} \left[ \Pi(\vec{u})(1 - \Pi(\vec{u}))^k + \Pi^k(\vec{u})(1 - \Pi(\vec{u})) \right] d\vec{u} \\ &= \frac{1}{(p+1)^n} \zeta_p^k(Y|Y). \end{aligned}$$

Thus,  $\zeta_p(Y, Z|Y, Z) > \frac{1}{(p+1)^{\frac{n}{p}}} \zeta_p(Y|Y)$ .

By Theorem 3.2.11, we can conclude that  $\frac{\zeta_p(Y, Z|X)}{\zeta_p(Y, Z|Y, Z)} < \frac{\zeta_p(Y|X)}{\zeta_p(Y|Y)}$ .  $\square$

Finally, we prove the next theorem which is a generalization of Theorem 1 in [7].

**Theorem 3.2.13.** *For any  $\epsilon > 0$ , there are absolutely continuous random vectors  $X$  and  $Y$  of arbitrary marginals but with the same dimension such that  $Y$  is completely dependent on  $X$  but  $\zeta_p(X|Y) \leq \epsilon$ .*

*Proof.* Since  $\zeta_p$  only depends on linkages, it is sufficient to prove the result for a uniform distribution.

First, we prove the result in the case of random variables. Let  $Y = kX \bmod 1$  where  $X$  has a uniform distribution. Then  $Y$  is also uniformly distributed. Clearly,  $\zeta_p(Y|X) = \zeta_p(Y|Y)$  since  $Y$  is a function of  $X$ . Also,

$$\begin{aligned} \frac{\partial}{\partial u} C_{Y,X}(u, v) &= \mathbb{P}(X \leq v | Y = u) \\ &= \begin{cases} \frac{1}{k} [kv - u] & \text{if } kv > u, \\ 0 & \text{otherwise} \end{cases} \end{aligned}$$

where  $\lfloor x \rfloor$  is the floor of  $x$ . Thus,  $\zeta_p^p(X|Y) \leq \frac{1}{k} \int |u|^p du$  and hence,

$$\begin{aligned} \zeta_p(X|Y) &\leq \left(\frac{1}{k}\right)^{\frac{1}{p}} \left(\frac{1}{p+1}\right)^{\frac{1}{p}} \\ &\leq \left(\frac{1}{k}\right)^{\frac{1}{p}}. \end{aligned}$$

Choosing  $k = \lfloor \epsilon^{-p} \rfloor + 1$ , we have  $\zeta_p(X|Y) \leq \epsilon$  as required.

In general, let  $(X_i, Y_i)$  where  $i = 1, \dots, n$  be independent copies of a random vector  $(X_0, Y_0)$  in which  $\zeta_p(Y_0|X_0) = \zeta_p(Y_0|Y_0)$  but  $\zeta_p(X_0|Y_0) \leq \frac{\epsilon}{n}$  and let  $X = (X_1, \dots, X_n)$  and  $Y = (Y_1, \dots, Y_n)$ . Then  $\zeta_p(Y|X) = \zeta_p(Y|Y)$  and

$$\begin{aligned} \zeta_p^p(X|Y) &= \int \left| \prod_{i=1}^n \frac{\partial}{\partial u_i} C_{Y_0, X_0}(u_i, v_i) - \prod_{i=1}^n v_i \right|^p du_1 \cdots du_n dv_1 \cdots dv_n \\ &\leq \sum_{i=1}^n n^{p-1} \int \left| \frac{\partial}{\partial u} C_{Y_0, X_0}(u, v) - v \right|^p du dv \\ &= n^p \zeta_p^p(X_0, |Y_0) \\ &= \epsilon^p \end{aligned}$$

and hence,  $\zeta_p(X|Y) \leq \epsilon$  as desired. □